

Conceptions of Assessment: Investigating What Assessment Means to Secondary and University Teachers

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Abstract

The overriding objective of this study was to investigate secondary and university teachers' assessment conceptions in an EFL context using a four-factor (*Student Accountability, School Accountability, Improvement, and Irrelevance*) teachers' conceptions of assessment (TCoA) inventory (Brown, 2006). Data of the study were collected by a questionnaire administered to secondary school ($n=336$) and university ($n=206$) teachers. Factor analyses (exploratory factor analysis (EFA), parallel analysis (principal component analysis (PCA)), dimension analysis (SPSS R-Menu v.2.0), and confirmatory factor analysis (CFA)) using SPSS v. 22.0 and AMOS v. 21.0 were examined. Results delineated a three-factor model (*Accountability, Improvement and Irrelevant*) with an endorsement of and a significant relationship between *Accountability* and *Improvement*. The ecological relationship between factors and indicators denoted teachers' misconceptions about assessment. Implications for future research on TCoA in similar-related contexts were also discussed.

Keywords: *Assessment literacy, TCoA, education policy, EFA, PCA, CFA, assessment misconceptions*

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Background to the Study

One of the significant roles of assessment conceptions is to preserve ethics and standards (Conley, 2005; Davies, 2008a). Such conceptions impact teachers' views on how to conceive of learning, teaching, and assessment intended to cope with instructional objectives and learning outcomes, a trend that was ignored during the early days when researchers (e.g., Lado, 1961) started theorizing about language testing. Many researchers (e.g., Inbar-Lourie, 2008; Malone, 2013; Stiggins, 2002; Stoyanoff & Chapelle, 2005) have accentuated varying definitions of assessment literacy. According to Taylor (2009), assessment literacy spans many stakeholders (Jeong, 2013) such as students, teachers, parents, policymakers, educators, and even governments. It is concerned with certifying learning (Brown, Kennedy, Fok, Chan & Yu, 2009; Kennedy, Chan & Fok, 2011), and it refers to teachers' theoretical and operational knowledge to design useful tests whose constructs are based on well-defined test specifications (specs). Further, assessment literacy is interpreted as the teachers' ability to use scores, analyse, comment, report results, make fair inferences on the test-takers' future (Qian, 2008) and write evaluation reports on instructional materials and language programs (Brown & Bailey, 2008; Davies, 2008a). In this regard, teachers are held to play a key role in assessment reforms.

Acquiring such literacy is embedded in a broader socio-cultural context where advocates of assessment literacy have addressed assessment conceptions (Brown, 2006; Brown et al., 2009; Gebril & Brown, 2013) and practices (Brown, 2002, 2004, 2008a, 2011). The challenging principle of assessment literacy (Malone, 2013) is to instill a new culture of assessment that straddles conceptions with practices and not disregard them. On the challenges of assessment literacy, Stiggins (2002, p. 762) contends that "few teachers are prepared to face the challenges of classroom assessment because they have not been given the opportunity to learn to do so." Fulcher and Bamford (1996) highlight some assessment principles at the core of any profession where tests are frequently used for examination or promotion. In addition, Inbar-Lourie (2008) stresses the deployment of multiple principles for assessment literacy, such as the disparity between formative and summative assessment, decision-making, classical vs. modern testing, different modes of assessment (i.e., alternative and traditional), the relevance of formative assessment to instruction, different measurement methods and the societal impacts of assessment practices. In this study, like the original TCoA inventory, the working definition of assessment literacy revolves around four major aspects: *student accountability*, *school accountability*, *improvement*, and *irrelevance*.

For Conley (2005), high-stakes assessment is potentially carried out for different purposes, the most important of which could be improvement and accountability. For instance, teachers use assessment to improve learning and teaching (Brindley, 2001; Scarino, 2013). Hamp-Lyons (1997) argues that the role of the tester is tied with accepting responsibility for all consequences of assessment conceptions and practices. Accountability is one aspect of the hallmark of assessment literacy, and it denotes

school and teacher responsibility for students' performance. Also, it has the purpose of justifying the worth of money spent on education. New funds are allocated based on students' achievement indicators. Conversely, if accountability is not preserved, assessment is then rendered void, inaccurate, and irrelevant. That is, despite the continually cited positive conceptions of assessment, such as student accountability, school accountability, and improvement along with its negative washback (Alderson & Wall, 1996), assessment can be viewed as irrelevant (Shohamy, 2001).

Diverse as the assessment purposes might appear, investigating all the different purposes is beyond the scope of this study. To mention a few, previous research has cogently investigated TCoA in different contexts, such as New Zealand (Brown, 2006, 2011), China, as an examination-driven context (Brown & Gao, 2015), Hong Kong (Brown et al., 2009), the Netherlands (Segers & Tillema, 2011) and Egypt (Gebril & Brown, 2013). Brookhart (2003) and Brown et al. (2009) elevated assessment as a way to improve teaching among New Zealand primary teachers and concluded that teachers had conflicting conceptions of assessment and that they endorsed *Accountability* and *Improvement*. Brown et al. (2009) instrumentally maintained that differences in assessment conceptions were significant while accentuating the role of assessment in improving and reinforcing learning. Consequently, they envisaged assessment reform in Hong Kong. In Iran, Brown, Pishghadam, and Sadafian (2012) underscored *Student Accountability* as being more relevant than *Improvement*. In their study on the effects of the high-stakes examination system among pre-and in-service Egyptian teachers, Gebril and Brown (2013) aptly pointed out that the TCoA inventory of New Zealand was inadmissible and found a three-factor model, instead of four, with a strong correlation between *Improvement* and *Student Accountability*. Brown (2011) stressed the importance of low-stakes assessment while drawing attention to the significant endorsement between the four main factors of the original TCoA (Brown, 2006). In the observed reviews of the literature on such assessment conceptions, it was found that most of these studies have significantly endorsed some factors at the expense of others, whether at the primary, secondary or tertiary level. Nonetheless, no study has investigated TCoA among university and secondary school teachers. This current study addresses this gap in the Tunisian context.

Assessment context in Tunisia

Even though they have their individual, societal, economic, political, and educational impingements, test impacts and uses have been overlooked in the Tunisian context. An overwhelming common reform trend induced the furtherance of a change approach by adhering to the implementation of a new education policy that has been seemingly imbued with a reconsideration of the teaching and assessment practices. Curriculum reform, perceived as tedious, has been open to continuous debate among many stakeholders. While a relentless policy of changing textbooks could hardly construe secondary education, tertiary education has bespoken the contrivance of a new educational system labeled *Bachelor*, *Masters*, and *Ph.D.* calling for curriculum

reform since 2005. Generally, assessment conceptions concatenated with the teachers' views of language and language learning have been marked by different backlogs. There has never been any formal initiative to investigate the Tunisian educational system. There is a dramatic lack of sound theoretical assessment knowledge among teachers, and thus, in principle, this lack has led to poor assessment quality. In this contrived situation, exams are perceived negatively by students, as evidenced by the lack of detailed and adequate assessment reports on their performance, apart from a score assigned out of 20. Even in professional events, such as conferences and workshops, there is hardly any organized event on assessment. It is also a neglected area at the MA and Ph.D. levels where researchers are not encouraged to investigate the prime considerations, foundations, conceptions, and practices of assessment.

Debating this worrying situation has yielded a counterbalancing effect policy seeking refuge in private tutoring that has started to be widespread even at the university level. Given the dramatic absence of stringent government guidelines that regulate assessment, a widespread public dissatisfaction with this educational system in general and language assessment, in particular, has sprung up. It is commonly envisaged that test contents in Tunisia most often answer and meet the teachers' agenda and expectations only, while the role of test-takers to show views of assessment, awareness, and even self-assessment has been marginalized. In such an academic milieu, students blame teachers for their wrong conceptions of assessment and their lack of transparency and professionalism and; thus, they revert to absenteeism. Graduates of English "learn" test design out of the experience as the scope of education has been flawed in preparing them to be good test designers. Towards such a propagated assessment policy, students, parents, and teachers have largely lost confidence in this system by labelling assessment irrelevant. Parents have cautioned against the teachers' policy, and policymakers have been aspiring for some improvements on the part of teachers. Caught in this vicious circle, no side has claimed responsibility for such assessment conceptions and practices. Along with these problems, other factors tie in with scoring tests that are predominantly subjective where specs are not properly delineated to the extent that sometimes test-takers fail to understand the basic requirements of how to process a test item. This is a low-stakes context where there is no plea for adherence to the implementation of international standardized tests.

The study

The purpose of this study was to investigate secondary and university TCoA and determine the relationship between the four major factors (*School Accountability, Student Accountability, Improvement and Irrelevant*) (Brown, 2006). While studies on TCoA are abundant (Brown, 2011; Gebril & Brown, 2013), hardly any study investigated such conceptions among EFL secondary and university teachers using Brown's inventory of TCoA (2006). Furthermore, in observing all the departments of English at the Tunisian tertiary level up to early 2015, there has been hardly any

mention of testing, evaluation, and assessment courses (Hidri, 2014), except for an introductory assessment course for the MA students for the year 2012 at the Faculty of Humanities and Social Sciences of Tunis, Tunisia. A major rationale behind this study was to raise assessment awareness among teachers, address the different assessment problems and suggest improvements accordingly. It follows then that the study aimed to answer the following research questions:

1. How do secondary and university teachers conceive of assessment?
2. What does CFA suggest concerning the TCoA inventory held by these teachers?
3. Is TCoA model as undertaken by Brown (2006) similar to the Tunisian TCoA model?

Participants, instruments, and procedure

Data of this study were gathered using a questionnaire on TCoA inventory (Brown, 2006) administered to university teachers ($n=206$) in 29 institutes and universities and secondary school teachers of English ($n=336$) in the 24 Tunisian governorates. Table 1 presents the demographic features of the participants. Data were collected through a period of nine months, from May 2014 to January 2015. The 336 sample, whose participation rate ranged from .6% to 14.9%, presented 69.9% of females and 30.1% males with 44.6 of teachers having a teaching experience that ranged from 11 to 15 years and .6% for teachers who had a teaching experience of ≥ 30 years with a mean of 3.00. For university teachers, 60.2% were females while 39.8%

Table 1 Demographic features of the participants ($n=542$)

| Features | Secondary ($n=336$) | | University ($n=206$) | |
|----------------------------|--------------------------|------|---------------------------|------|
| <i>Gender</i> | Total | % | Total | % |
| Female | 235 | 69.9 | 124 | 60.2 |
| Male | 101 | 30.1 | 82 | 39.8 |
| <i>Teaching Experience</i> | | | | |
| 1-5 | 48 | 14.3 | 61 | 29.6 |
| 6-10 | 56 | 16.7 | 54 | 26.2 |
| 11-15 | 150 | 44.6 | 49 | 23.8 |
| 16-20 | 24 | 7.1 | 20 | 9.7 |
| 21-25 | 47 | 14.0 | 12 | 5.8 |
| 26-30 | 9 | 2.7 | 6 | 2.9 |
| More than 30 | 2 | .6 | 4 | 1.9 |
| Mean | | 3.00 | | 2.52 |

were males with 29.6% of a teaching experience that ranged from 1 to 5 years. The least percentage was among teachers who had an experience of ≥ 30 years with 1.9%.

The mean of teaching experience is 2.52, and the participation rate ranged from .5% to 23.3%.

The TCoA inventory (Brown, 2006) contains four factors with 27 items or indicators (Appendix A): *School Accountability* and *Student Accountability* as first-order factors, *Improvement*, and *Irrelevance*, as second-order factors. *School Accountability* and *Student Accountability* include three first-order contributing indicators each: assessment **provides** information on schools, assessment is **accurate**, assessment **evaluates** schools and assessment **categorises** students, assessment **assigns** scores to students' work, and assessment **determines** students' qualifications, respectively. *Improvement* includes four second-order contributing factors, each of which has three other contributing indicators: 1) assessment **describes** abilities (assessment determines the quantity of learning, establishes learning content, and measures meta-cognitive thinking skills among students) 2) assessment **improves learning** (assessment provides feedback on students' performance, feeds back learning needs to students and improves students' learning), 3) **improves teaching** (assessment is integrated with teaching, modifies teaching and allows different instructions for students) and 4) assessment is **valid**. The fourth factor, *Irrelevance*, has also three second-order contributing factors: 1) assessment is **bad** (against beliefs, unfair, interferes with teaching), 2) assessment is **ignored** (little use of results, results are filed and ignored and impacts teaching) and 3) assessment is **inaccurate** (measurement error, error and imprecision, and imprecise process). The scale used in the TCoA inventory was a five-point agreement scale of *strongly disagree*, *disagree*, *agree*, and *strongly agree* with a mid-position of "*undecided*". TCoA inventory was administered online to 542 participants. Since the TCoA inventory was administered online, where the respondents had to select one option from each row to proceed, there were not any missing or invalid cases. Cronbach alpha for the agreement Likert scale was $\alpha = .78$ ($M = 82.86$ and $SD = 13.983$) for all the 27 item-data sets.

Data Analyses

Given the complexity of factor analysis, data analyses were carried out in four phases (Table 2). The types of analyses were: EFA, parallel analysis (Monte Carlo PCA), dimension analysis, and CFA, which were geared towards defining the appropriate number of factors and checking whether the data fitted Brown's original inventory of TCoA (2006). EFA relied on descriptives of factor extraction, selection, and rotation with scree plot and goodness-of-fit-test. Monte Carlo PCA and scree plot stressed the use of eigenvalues, mean and percentile random data eigenvalues, dimension analysis used the SPSS R-Menu v2.0 to get the exact number of factors, and CFA was concerned with fit indices, using AMOS v 22.0 so that the data would fit the model.

Phase one: Exploratory Factor Analysis

Using SPSS 22.0, EFA was used as a precursor to CFA to investigate the possible factor structure. In processing EFA (Table 2, phase 1), five statistics were used:

Table 2 Data collection and analyses

| Phases | Tests | Results | | | Rationale |
|-------------|--|--|--|--|---|
| Phase one | <i>EFA:</i> - Descriptives: Kaiser-Meyer Olkin (KMO) - Bartlett's Test of Sphericity - Factor extraction (ML), - Factor selection (Eigenvalues and scree plot, and goodness-of-fit-test) - Factor rotation (Promax) | KMO: .794 χ^2 : 2487.726 <i>df</i> : 351 $p = .000$ 8 factors χ^2 : 276.832, <i>df</i> : 163, $p = .000$ | KMO: .782 χ^2 : 2487.726 <i>df</i> : 253 $p = .000$ 6 factors χ^2 : 226.130, <i>df</i> : 130, $p = .000$ | | - To get the possible factor structure - To define the items whose values are $\geq .30$ - To check if the ratio of χ^2/df are significant at $p = .000$ |
| Phase two | - <i>Parallel analysis:</i> Monte Carlo PCA - Scree plot | 6 factors 6 factors | | | - To check the number of factors that had to fit the model. |
| Phase three | - <i>Dimension analysis:</i> SPSS R-Menu v.2.0. - RMSR eigenvalues - Parallel analysis - Goodness-of-fit-test | 3 to 8 factors Eigenvalues > mean = ($n = 8$), parallel analysis ($n = 6$), optimal coordinates ($n = 3$) χ^2 : 1027.58, <i>df</i> : 273, $p = .000$ | | | - To check the number of factors yielded by parallel analysis |
| Phase four | <i>CFA:</i> AMOS 21.0 - χ^2 , <i>df</i> , χ^2/df , SRMR, RMSEA, CFI and TLI Reliability analysis: Cronbach Alpha Pearson Correlation analysis | Original model (Appendix B, 27 items) - χ^2 : 1805.788 - <i>df</i> : 312 - χ^2/df : 5.788 - SRMR: .17 - RMSEA: .094 - CFI: .52 - TLI: .46 - $p = .000$ - $\alpha = .78$ ($M = 82.86$, $SD = 13.983$) | 6 factors (Appendix C, 22 items) - χ^2 : 525.278 - <i>df</i> : 194 - χ^2/df : 2.708 - SRMR: .10 - RMSEA: .056 - CFI: .84 - TLI: .80 - $p = .000$ - $\alpha = .75$ ($M = 70.15$, $SD = 12.087$) | 3 factors (Tunisian TCoA (Figure 5) (15 items) - χ^2 : 266.872 - <i>df</i> : 87 - χ^2/df : 3.067 - SRMR: .089 - RMSEA: .062 - CFI: .90 - TLI: .88 - $p = .000$ - $\alpha = .78$ ($M = 56.93$, $SD = 9.68$) | - To check if the fit value is acceptable - To check the goodness-of-fit indices (since relying on χ^2 might not produce valid fit indices). |

descriptives (KMO), Bartlett's Test of Sphericity, factor extraction (Maximum Likelihood (ML)), factor selection (eigenvalues and scree plot), and factor rotation loadings (goodness-of-fit-test and Promax). Factor extraction and factor rotation using ML were meant to define the exact number of factor loadings to fit the model. To produce high loadings whose values should be $\geq .30$, the pattern matrix was considered to investigate the loadings resulting from the appropriate relationship between factors and indicators. Therefore, the indicators that were $\leq .30$ or the ones which were linked to more than one factor were deselected.

Phase two: Parallel analysis

After the rotation phase, a parallel analysis was used (see Hayton, Allen, & Scarpello, 2004 on parallel analysis). Specifically, the Monte Carlo PCA parallel analysis (e.g., scree plot) was carried out to determine the statistically significant eigenvalues based on random data generation. The sample data was set at 100 cases to produce accurate analysis. The desired percentile was set at 95%, then a PCA was set at 2, and the data generation parallel analysis was set at 1. These tests were implemented to determine the factor structure.

Phase three: Dimension analysis

Dimension analysis of the data using the SPSS R-Menu v2.0 was conducted to opt for “statistical and graphical computing” (Courtney, 2013, p. 5) to estimate and reach the right number of factors. Unlike EFA, dimension analysis produced different results from the ones of phases one and two.

Phase four: Confirmatory factor analysis

Extant research has shown that CFA, pertained to Structural Equating Model, has been widely used to investigate the relationship paths between variables, the correlation between observed and latent variables (Miller, Davidson, Schindler, & Messier, 2013), to test data fit (Brigman, Wells, Webb, Villares, Carey & Harrington, 2015) and to check indicators' influence on factors (Hu & Bentler, 1998, 1999). CFA was carried out using AMOS to check whether the data would yield another Tunisian inventory since the original model was found to be inadmissible. Many researchers have maintained that a sample of ≤ 400 respondents is appropriate to claim the validity of CFA results (Brown, 2006). CFA includes the following criteria: chi-square statistic (χ^2), degree of freedom (df), and the ratio of chi-square to degrees of freedom (χ^2/df). To counterbalance chi-square sensitivity, four indices were considered: Root mean squared error of approximation (RMSEA), root mean square residual (RMSR), Tucker-Lewis non-normed fit index (TLI), and comparative fit index (CFI).

Values of CFA range from .0 to 1 (Brown, 2006; Cokley, 2014). In addressing the latent trait modeling, Brown (2006), Browne and Cudeck (1992) and Hu and Bentler (1999) suggested the following good fit indices values: CFI $\leq .95$, RMSEA $\approx .06$. For other researchers (e.g., Hair, Tatham, Anderson & Black, 2005), the RMSEA value

of 0.10 is indicative of unacceptable fit, 0.08-0.10, mediocre fit, 0.06-0.08, acceptable fit, and 0.01-0.06, close fit. In all cases, it should be $\leq .70$. The χ^2/df ratio should range from 2 to 3 as indicative of good fit (Choo, Walsh & Teyl, 2013), while values of CFI and TLI should be 0.95 as good fit and 0.90 as acceptable fit. For SRMR, values of 0.08 indicate good fit, and 0.12 would indicate acceptable fit. The goodness-of-fit indices were considered based on Brown's (2006) and Hu and Bentler's (1998, 1999).

Results

EFA results

All the 27 items of the original TCoA were considered to yield a possible factor structure of the appropriate latent variables and indicators with high loadings and, therefore, check whether this inventory was admissible to the Tunisian context. Like the original model (Brown, 2006), six first-order items (Appendix B) were loaded onto *School Accountability* and *Student Accountability*, respectively. Four second-order items with three items each were loaded onto *Improvement*. As for the fourth factor (*Irrelevance*), three items included three contributing factors each, respectively. At this level, the model was found to be inadmissible. Therefore, EFA was carried out to define the exact number of factors to yield appropriate fit indices. For instance, EFA using parallel analysis, pattern matrix, and scree tests suggested an eight-factor model. Loadings of $\leq .30$ were discarded. The rotated factors using ML were estimated at an eigenvalue of $\geq .30$. The KMO was .794, indicating an acceptable

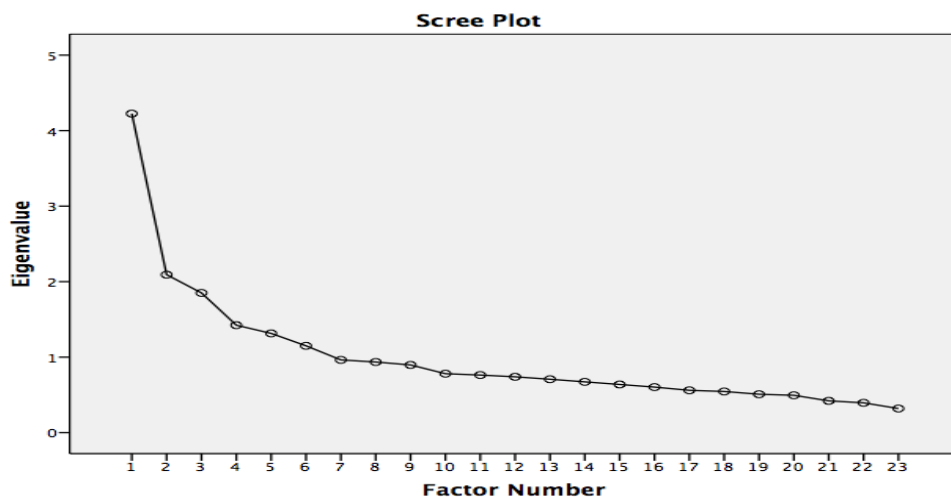


Figure 1. Scree plot of the eigenvalues (EFA, 8 factors)

fit value of all the data (Table 2). The Bartlett's test of sphericity χ^2 was 2487.726 with a df of 351. The goodness-of-fit-test of χ^2 was 276.832 with a df of 163 was significant at .000, as expected because of the huge dataset ($n=542$). Based on the eigenvalues of the 27 indicators, Figure 1 shows 8 factors that the data set produced. The number of factors was determined once the slope of the figure changed. The number of factors

was determined by counting all the instances that occurred before this slope. However, Brown (2006) claims that the results of the scree plot test may sometimes be ambiguous.

Table 3 Raw data eigenvalues, mean and percentile random data eigenvalues

| Root | Raw Data | Means | Prcntyle |
|-----------|----------|----------|----------|
| 1.000000 | 4.764549 | 1.436147 | 1.491098 |
| 2.000000 | 2.370154 | 1.372303 | 1.424023 |
| 3.000000 | 2.006776 | 1.324681 | 1.363655 |
| 4.000000 | 1.526081 | 1.282012 | 1.320723 |
| 5.000000 | 1.398338 | 1.241457 | 1.270199 |
| 6.000000 | 1.265056 | 1.206071 | 1.231711 |
| 7.000000 | 1.148393 | 1.174377 | 1.196837 |
| 8.000000 | 1.102225 | 1.146129 | 1.168262 |
| 9.000000 | .990467 | 1.117137 | 1.145560 |
| 10.000000 | .919678 | 1.089348 | 1.111021 |

Analysis of the pattern matrix yielded unfit values in the dataset at the initial phase with some outliers. To remedy this, the items that did not load anywhere or those loaded with two factors were deselected from the pattern matrix analysis, such as items 2, 3, 10, 12, and 22. Therefore, the pattern matrix yielded loadings ranging from .334 (indicator 9) to .935 (indicator 20). The KMO measure of sampling adequacy was .782, and χ^2 was 2487.726, with a *df* of 253 was still significant at .000. The goodness-of-fit-test for the chi-square was 226.130 with *df* of 130, significant at .000. EFA results indicated the loadings of six factors along with their commonalities: E.g., 3 (.756), 4 (.703), 14 (.632), 5 (.452), 13 (.431) and 1 (.423) loaded on factor 1; indicators 20 (.935), 21 (.656), and 19 (.576) loaded on factor 2, indicators 27 (.607), 11 (.547), 6 (.506), 16 (.472) and 15 (.412) loaded on factor 3, indicators 24 (.683), 25 (.549) and 23 (.382) loaded on factor 4, indicators 17 (.659), 18 (.451) and 26 (.357) loaded on factor 5 and indicators 8 (.600), 7 (.497) and 9 (.334) loaded on factor 6. Still, at this level, the factors had conflicting indicators. To solve this problem, parallel analysis was implemented.

Parallel results

Parallel analysis was utilized to check further the number of factors that had to fit the model. Table 3 indicates the results of the Monte Carlo PCA analysis that yielded six factors. It presents raw data eigenvalues, mean and percentile random data eigenvalues carried out on the 27 items for all the participants (N= 542), 10 items of the framework, and 100 random data sets were generated and a percentile of 95. Column two, raw data, calculated the PCA of eigenvalues on the correlation matrix that matched the actual data of the SPSS. The first item has a component eigenvalue of 4.764549, the mean (which was anchored at 50%) is 1.436147, and a percentile value of 1.491098 anchored at 95 up to item 6 with an eigenvalue of 1.265056, mean of

1.206071, and a percentile of 1.231711. The percentile has to be larger than the mean of eigenvalues, and it, therefore, indicates a factor. The eigenvalues of the first six raw data values were statistically significant since their values were larger than the benchmark criterion value of the percentile. Based on this analysis, data analysis

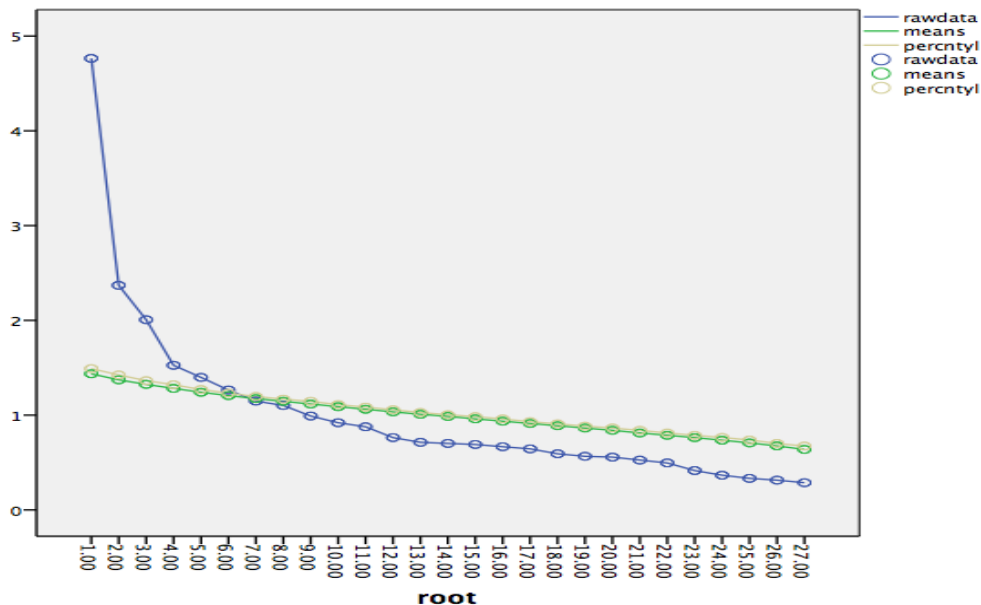


Figure 2. Data sequence plot (Monte Carlo PCA)

pointed to six-factor components. A graphical presentation of the Monte Carlo PCA using scree plot (Figure 2) shows six factors. The (blue) line with the vertical orientation presented the raw data, the green line indicated the means, and the last one (brown line) showed the eigenvalues and it, therefore, showed a factor. The eigenvalues of the first six raw data values were statistically significant since their values were larger than the benchmark criterion value of the percentile. Based on this analysis, data of the study pointed to six-factor components to extract from this analysis. The two lines (brown and green) intersected with the scree, vertical line, and factors above the competing eigenvalue lines represented the number of factors that should be extracted. The number of factors was selected based on the intersection lines. Therefore, the points above the intersection lines represented the number of factors, which showed six factors.

Dimension analysis results

In the dimension analysis results, a correlation matrix of heterogeneous (two steps) analysis indicated the retention of 3 factors (Figure 3). However, the analysis of comparison data (fit-to-comparison data and Pearson analysis) (Figure 4) indicated the retention of 8 factors. Results of dimension analyses using the SPSS R-Menu v2.0

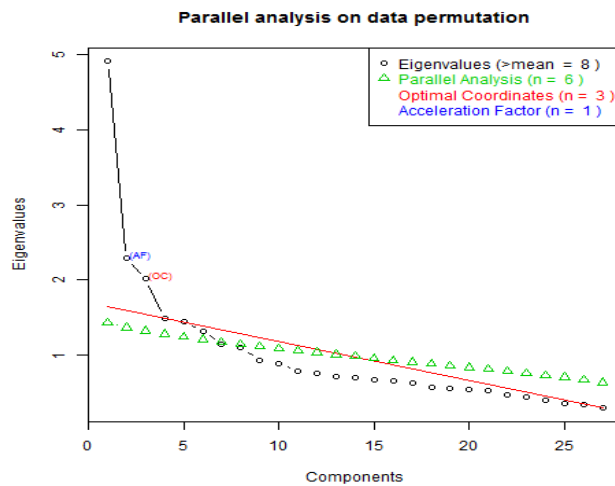


Figure 3. Parallel analysis on data permutation

Table 4 Factor correlation matrix

| Factor | 1 | 2 | 3 |
|--------|-------|-------|-------|
| 1 | 1.000 | .187 | -.240 |
| 2 | .187 | 1.000 | -.118 |
| 3 | -.240 | -.118 | 1.000 |

Extraction Method: Maximum Likelihood.

Rotation Method: Oblimin with Kaiser Normalization.

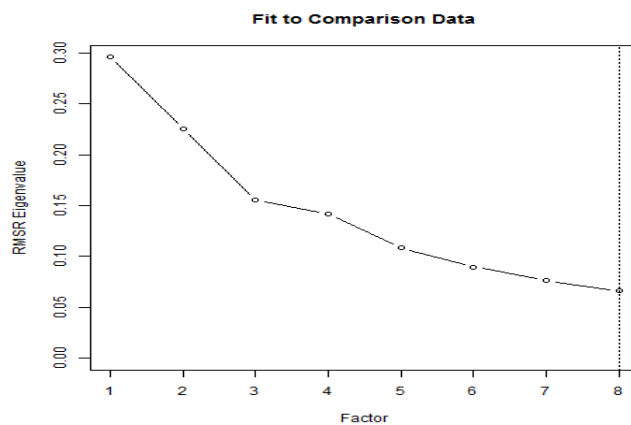


Figure 4. Fit to comparison data (8 factors)

Table 5 Fit to comparison data of 8 factors

| Nb. of factors | RMSR Eigenvalue | p-value |
|----------------|-----------------|---------|
| 1 factor | .296 | NA |
| 2 factor | .225 | .000 |
| 3 factor | .156 | .000 |
| 4 factor | .142 | .000 |
| 5 factor | .108 | .000 |
| 6 factor | .090 | .000 |
| 7 factor | .076 | .000 |
| 8 factor | .066 | .000 |

(Courtney, 2013) showed a data set that ranged from 3 (Figure 3) to 8 factors (Figure 4), and the factor correlation matrix (Table 4) showed 3 factors; however, the fit to comparison data eigenvalues indicated 8 factors (Table 4).

CFA results

Once the number of factors was determined, CFA was used. Data of the study on the TCoA in Tunisia indicated a range of 3 to 8 factors. The initial phase of the study constituted the use of CFA. Recall that AMOS covariance matrices analyses between factors demonstrated that Brown's model (2006) was found to be inadmissible (Appendix B) with the following fit indices: $\chi^2 = 1805.788$, $df = 312$, $\chi^2/df = 5.788$, CFI = .52, TLI = .46, and RMSEA = .094 and SRMR = .17. The loadings from *Improvement to describe* and *improves learning* were beyond the range as they both had a value of 1.15 and 1.02 and 1.39, respectively. Recall that in CFA, such values should range from .0 to 1. Also, the covariance matrices between *Improvement* and *Student Accountability* on

the one hand and between *School Accountability* and *Irrelevance* on the other were beyond the range with values of 1.15 and 1.53, respectively.

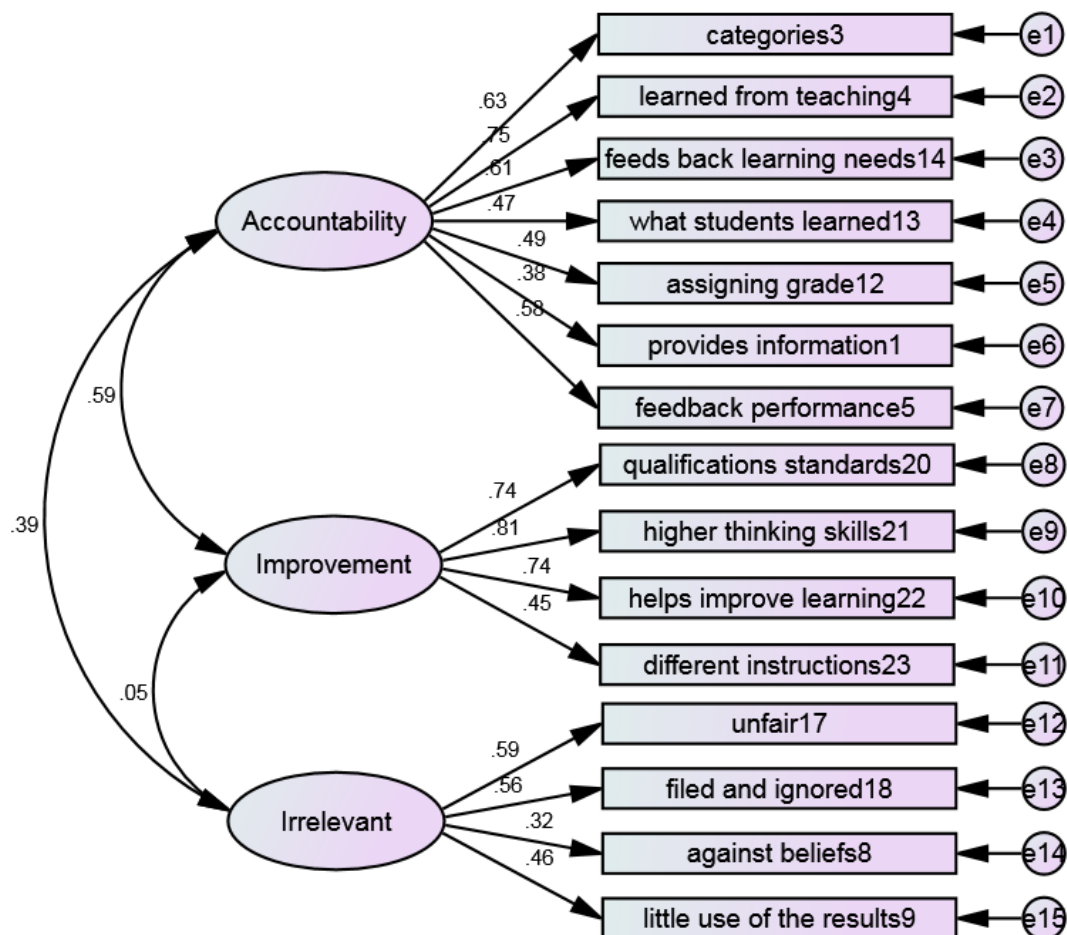


Figure 5. The Tunisian model of TCoA: χ^2 : 266.872, df : 87. χ^2/df : 3.067, SRMR: .89, CFI: .90, TLI: .88, RMSEA: .062. All factor loadings were significant at $p = .000$ level.

After deselecting indicators 2, 3, 10, 12 and 22, the model yielded 6 factors with acceptable fit indices (see Appendix C): $p = .000$, χ^2 : 525.278, df : 194. χ^2/df : 2.708, SRMR: .10, CFI: .84, TLI: .80, RMSEA: .056. However, even though the model had high loadings, it nonetheless was inadmissible, since many indicators did not load onto the appropriate factors, such as 19, 27, 26 and 7. In trying to attach the indicators with the appropriate factors, the preliminary Tunisian model (Appendix C) was edited and the deselected factors were reintegrated in CFA to recheck the factor structure. Figure 5 presents the following three-factor model with first order factors attached to 15 indicators. Factor 1, *Accountability*, included indicators and loadings as follows: 3 ($r = .63$), 4 ($r = .75$), 14 ($r = .61$), 13 ($r = .47$), 12 ($r = .49$), 1 ($r = .38$) and 5 ($r = .58$). Factor 2, *Improvement*, included indicators 20 ($r = .74$), 21 ($r = .81$), 22 ($r = .74$) and 23 ($r = .45$) and factor 3, *Irrelevant*, included indicators 17 ($r = .59$), 18 ($r = .56$), 8 ($r = .32$) and 9 ($r = .46$).

The fit indices of the Tunisian model are the following: $p = .000$, χ^2 : 266.872, df : 87. χ^2/df : 3.067, SRMR: .89, CFI: .90, TLI: .88, RMSEA: .062 and, therefore, a three-factor trimmed model version was found to be admissible. It could be cautiously argued that this model somehow yielded acceptable fit indices.

In addressing the internal consistency of the factors, the reliability coefficients indices using Cronbach Alpha (α) were considered for both populations: secondary school and university teachers. Reliability of the three factors was calculated among both populations. Results of *Accountability* among secondary school teachers indicated $\alpha = .81$ ($M = 23.05$, $SD = 5.984$), *Improvement* $\alpha = .78$ ($M = 12.96$, $SD = 3.738$) and *Irrelevant* $\alpha = .63$ ($M = 12.60$, $SD = 3.730$). As for university teachers' data, *Accountability* $\alpha = .60$ ($M = 20.67$, $SD = 5.045$), *Improvement* $\alpha = .76$ ($M = 13.16$, $SD = 3.847$) and *Irrelevant* $\alpha = .16$ ($M = 10.33$, $SD = 2.907$). The entire scale for both populations yielded the following results: $\alpha = .75$ for *Accountability* of 7 items ($M = 22.15$, $SD = 5.758$), $\alpha = .77$ for *Improvement* of 4 items ($M = 13.04$, $SD = 3.778$), $\alpha = .54$ for *Irrelevant* of 4 items ($M = 11.74$, $SD = 3.610$) and $\alpha = .78$ for the entire scale statistics of 15 items ($M = 56.93$, $SD = 9.68$).

In addition, a Pearson correlation coefficient analysis was conducted to investigate the correlation patterns between the 15 indicators. For results on secondary school teachers' data, the *Accountability* correlation between *learned from teaching* and *feedback performance*, with coefficients of .55, was significant at ($p \leq .001$). The least significant correlation between *provides information* and *feedback performance* was .28 ($p \leq .001$). For *Improvement*, correlation coefficients between *higher thinking skills* and *qualification standards* were .63 ($p \leq .001$), while they were significant at .31 ($p \leq .001$) between *different instructions* and *higher thinking skills*. For *Irrelevant*, correlation coefficients between *little use of results* and *against beliefs* were .41 ($p \leq .001$); however, they were .15 ($p \leq .001$) between *against beliefs* and *filed and ignored*. For *Accountability* results among university teachers, the Pearson correlation coefficients between *categories* and *learned from teaching* were .40 ($p \leq .005$) and .19 ($p \leq .005$) between *what students learned* and *feeds back learning needs*, while the coefficients were .14 ($p \leq .001$) between *assigning a grade* and *categories*. As for *Improvement*, correlation coefficients between *higher thinking skills* and *qualifications standards* were .62 ($p \leq .001$) and .25 ($p \leq .001$) between *higher thinking skills* and *different instructions*. As for *Irrelevant*, correlation coefficients between *unfair* and *filed and ignored* were .33 ($p \leq .001$). However, the Pearson correlation coefficients were not significant since they yielded the following: -.072 between *filed and ignored* and *against beliefs*. In observing the model, the *Accountability* correlation between *learned from teaching* and *categories* was .49 ($p \leq .001$), while the correlation between *feeds back learning needs* and *provides information* was .17 ($p \leq .001$). As for *Improvement*, the Pearson correlation coefficients between *higher thinking skills* and *qualifications standards* were .63 ($p \leq .001$) but were .27 ($p \leq .001$) between *different instructions* and *higher thinking skills*. As for *Irrelevant*, the Pearson correlation coefficients between *unfair* and *filed and ignored* were .36 ($p \leq .005$), while they were .20 ($p \leq .005$) between *unfair* and *against beliefs* were but .09 ($p \leq .001$) between *filed and ignored* and *against beliefs*.

Discussion

The study investigated secondary and university TCoA in an EFL context. To approach this, different analyses were adopted. A conspicuous result of the study was levelled at the wrong and conflicting assessment conceptions among Tunisian teachers of English. The latent structural paths between variables and indicators of the initial model were different from previously observed models (e.g., Brown, 2006; Brown & Goa, 2015; Gebril & Brown, 2013) and, therefore, it resulted in a poor fit. This is reflected in the beyond-range values of the relationship (See Appendix B for covariance matrices). However, based on the different types of analyses, one may provide a baseline that the data collected on Tunisian secondary and university teachers partially fitted the model, resulting in high loadings and different configurations of the relationship between factors and indicators. This partial fit indicated that assessment conceptions were held to be divergent among teachers of English.

While previous studies highlighted the results of 27 items with appropriate loadings but with different combinations of factors and indicators (e.g., Brown & Michaelides, 2011), this study had different results in that the 15 indicators loaded onto different factors. Structurally speaking, the paths among factors and between factors and indicators (Appendix C) were disparate and divergent from the original TCoA (indicators 19, 27, 26, and 7 are a case in point) even though they yielded high loadings. However, like other studies (e.g., Kitiashvili, 2014), a central axiom here lies at the heart of the conjointly problematic assessment misconceptions among secondary and university teachers despite the correlation of high loadings. This had to be expected given the stakeholders' mundane attitudes towards assessment where assessment courses or professional development events are not attributed their due relevance. Thus, this lack of assessment expertise does not bode well for a clear and sustained assessment policy. An essential comment of this discussion that can be upheld is that considering relevant assessment literacy needs not be ignored or taken for granted. Instead, it should be based on objective and well-sustained guidelines where teachers are expected to hold a key role. The conflicting conceptions of assessment might also impact teachers' practices. This idea was echoed in the study conducted by Cheng, Rogers, and Hu (2004).

The Tunisian model of TCoA was divergent from previous studies' models, such as Brown and Michaelides (2011) and Brown (2011). Even though the Tunisian model might look similar to the simplified New Zealand model, it is nevertheless different since it resulted in a trimmed number of factors and indicators, 15 instead of 27 (see Appendix D on the number of deselected items from the original TCoA). The Tunisian model of TCoA had direct relationships between indicators and their factors. Like previous studies (e.g., Gebril & Brown, 2013), this study showed a strong correlation between *Accountability* and *Improvement*. What could be deduced is that even though assessment has not been attributed its role in such context, teachers still conceive of assessment positively. This result is echoed in other studies (e.g., Brown, 2006). It

could also be concluded that secondary and university TCoA were enmeshed in disparate attitudes in a sense that while teachers conceived of assessment as a way to improve learning, they also perceived it as irrelevant. In the Tunisian context, teachers have been held responsible for the testing quality on the ground that teaching and testing do not prepare graduates to be operational in their field of work. In this regard, students' criticism has been lodged at teachers whose exams have been perceived as a heavy burden. Meaning-ladenness of assessment has been suffused with vagueness and disparity given the dramatic lack in theoretical underpinnings and practical tips of assessment, and unfortunately, students graduate with no basic knowledge or training in assessment. Notwithstanding the *ad hoc* leaps in reform that have marked education in Tunisia, still, the assessment policy is vague, as teachers who do not have clear assessment visions are likely to miss the learning objectives and, therefore, make testing an irrelevant task. Other stakeholders such as learners, parents, and policymakers should attempt to take cognizance of the teachers' divergent educational backgrounds. This is manifested in the wrong loadings of indicators onto the latent variables, which could be true due to the misspecifications of the TCoA.

Unlike the previous educational system that rested on summative assessment, newer developments within the current assessment policy in Tunisia proffered the idea of a revamp of formative assessment. Unfortunately, unlike other studies (e.g., Brindley, 2001), assessment in Tunisia has been perceived to evaluate students on whether to pass or fail and not to develop their critical thinking skills to overcome the different language problem-solving tasks tests may contain. To strive against the teachers' bias for summative assessment, many stakeholders, such as policymakers, should consider the implementation of both formative and summative assessment. Whether at school or university, one result of this study showed that assessment was generally used in a summative way and not as a diagnostic tool to establish a comprehensive view of the students' language ability. Item writers overlook test specs on the ground that they are experienced enough to design fair tests. One trenchant criticism at this level is that test designers write items that measure a limited range of skills and sub-skills or tests that do not measure what they are intended to measure. This reflected the teachers' lack of sound knowledge about language assessment.

Limitations, implications, and recommendations

Results of the study indicated that much research is needed to investigate the TCoA at the different educational levels in Tunisia. Analysing data of the study on secondary and university teachers separately could have led to other results. This posed some limitations since the implementation of factor analysis necessitates the use of ≥ 400 participants. Investigating a comparison of the model between the two populations might enlighten the researcher to compare and contrast assessment conceptions in the two models. In addition, using mixed-methods research, such as interviews and item analysis of test scores, could undoubtedly provide more insights into the teachers'

assessment conceptions that would have direct implications for teaching and assessment.

In addressing the implications of the study, it could be argued that, as part of methodological implications, the use of factor analysis (EFA, PCA, SPSS R v.2.0, and CFA) helped to unveil assessment conceptions and to address the factor structure of acceptable or good fit models. In the pedagogical implications, this study presented practical steps in investigating teachers' beliefs about assessment, which are directly linked to classroom practices. Like other studies (Gebril & Brown, 2013), the current study might plausibly be relevant to such contexts in probing into the significant assessment requirements as well as teaching conceptions. As for the research implications, addressing such assessment enterprises was intertwined with some wrong assessment beliefs and views that can potentially affect assessment practices. Such practices are upheld by many teachers who are encouraged to invest a lot in formative assessment to provide test-takers with continuous feedback on their performance. Further research on the teachers' conceptions of assessment is needed in the Tunisian context. An exhortation to teachers is to team up, consider test specs, pilot test items, administer tests, analyse test scores and reconsider test specs where collaborative initiatives should be highlighted to improve the assessment quality. In line with the different calls for change, all stakeholders should seize this opportunity and consider all challenges pertained to assessment. It is highly recommended that universities in Tunisia reconsider their curricula to stress the relevance of administering assessment courses for students before graduation.

One of the study outcomes was the call for MA and Ph.D. students, in particular, and researchers, in general, to re-investigate the teachers' assessment conceptions and, by extension, practices. The adherence to a code of practice among all stakeholders is a diligent necessity. In addition, implementing high-stakes professional assessment standards should have pertained to the learners' knowledge of instructional objectives and testing outcomes. In this regard, test designers should be aware of the different testing methods, such as classical, modern, discrete-point, integrative or communicative, since this awareness reflects their views of language and language learning. Another challenge that faces test designers is the dilemma of whether the training in testing, if any, can enable them to design useful tests. To develop their assessment literacy, teachers should consider the following: test specs, assessment ethics, professional assessment standards, testing outcomes, the relevance of placement tests, alternative forms of assessment, setting up exam boards, needs analysis of students' assessment lacks, the relevance of using international standardized exams as entry and/or exit exams, item analysis, testing and curriculum design, critical approaches to assessment, program evaluation, and classroom testing practices. Since assessment is the backbone of any educational system, assessment courses should be considered at the university level. As for secondary school teachers, supervisors of English themselves have to target assessment literacy in their regular training sessions. Whether at the secondary or university level, assessment in Tunisia has to be taken seriously and meaningfully. However, it is still norm-referenced, and it has been widely

implemented at the expense of criterion-referenced assessment, and probably teachers, test designers, and other stakeholders have not been aware of this dichotomy.

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Appendix A Teachers' Conceptions of Assessment (Brown, 2006, p. 168)

Factors and indicators

Assessment Makes Schools Accountable

Assessment provides information on how well schools are doing.

Assessment is an accurate indicator of a school's quality.

Assessment is a good way to evaluate a school.

Assessment Makes Students Accountable

Assessment places students into categories.

Assessment is assigning a grade or level to student work.

Assessment determines if students meet qualifications standards.

Assessment Improves Education

Assessment Describes Abilities

Assessment is a way to determine how much students have learned from teaching.

Assessment establishes what students have learned.

Assessment measures students' higher-order thinking skills.

Assessment Improves Learning

Assessment provides feedback to students about their performance.

Assessment feeds back to students their learning needs.

Assessment helps students improve their learning.

Assessment Improves Teaching

Assessment is integrated with teaching practice.

Assessment information modifies on-going teaching of students.

Assessment allows different students to get different instruction.

Assessment is Valid.

Assessment results are trustworthy.

Assessment results are consistent.

Assessment results can be depended on.

Assessment is Irrelevant

Assessment is Bad

Assessment forces teachers to teach in a way against their beliefs.

Assessment is unfair to students.

Assessment interferes with teaching.

Assessment is ignored

Teachers conduct assessments but make little use of the results.

Assessment results are filed and ignored.

Assessment has little impact on teaching.

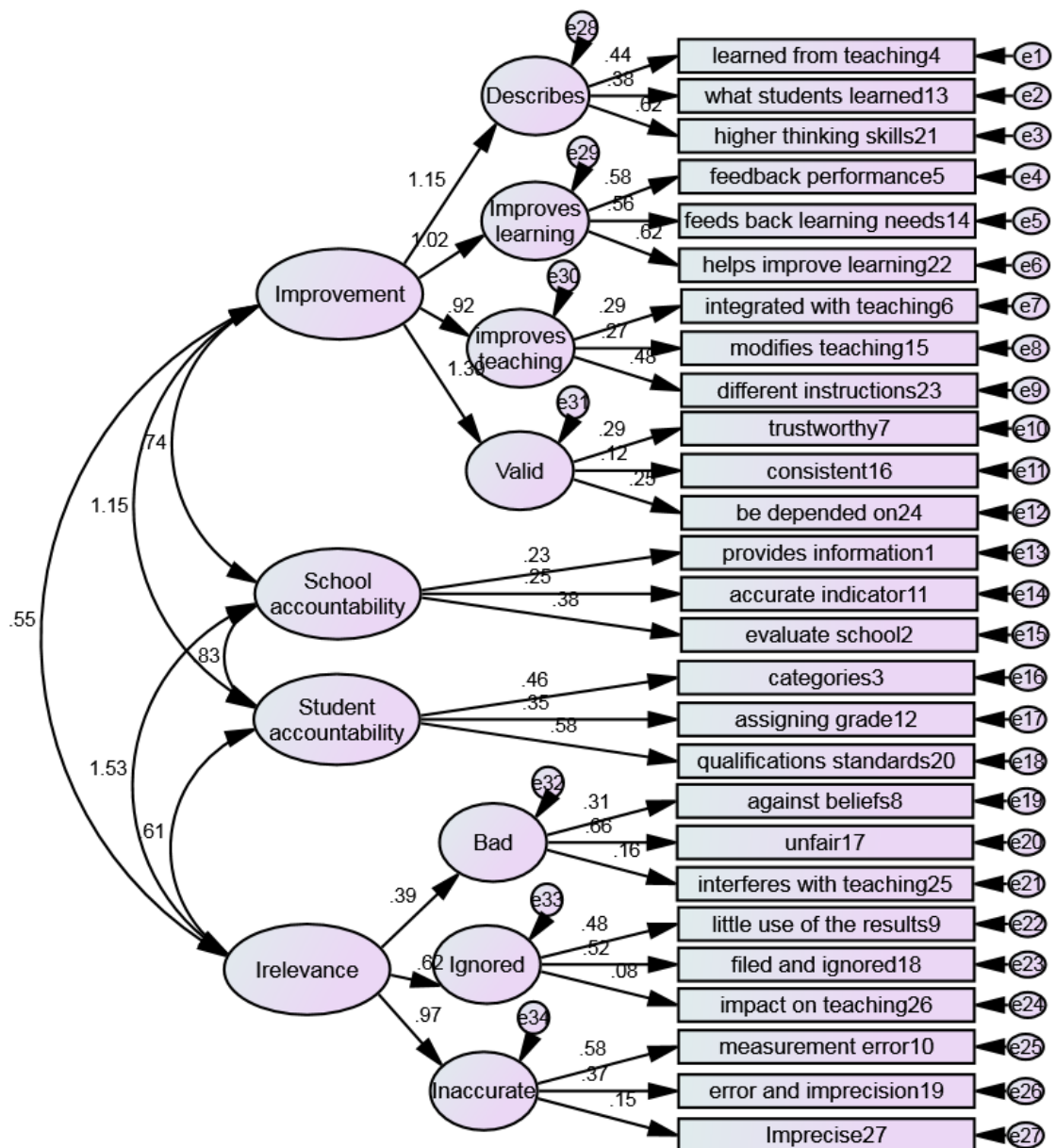
Assessment is Inaccurate

Assessment results should be treated cautiously given measurement error.

Teachers should take into account the error and imprecision in all assessment.

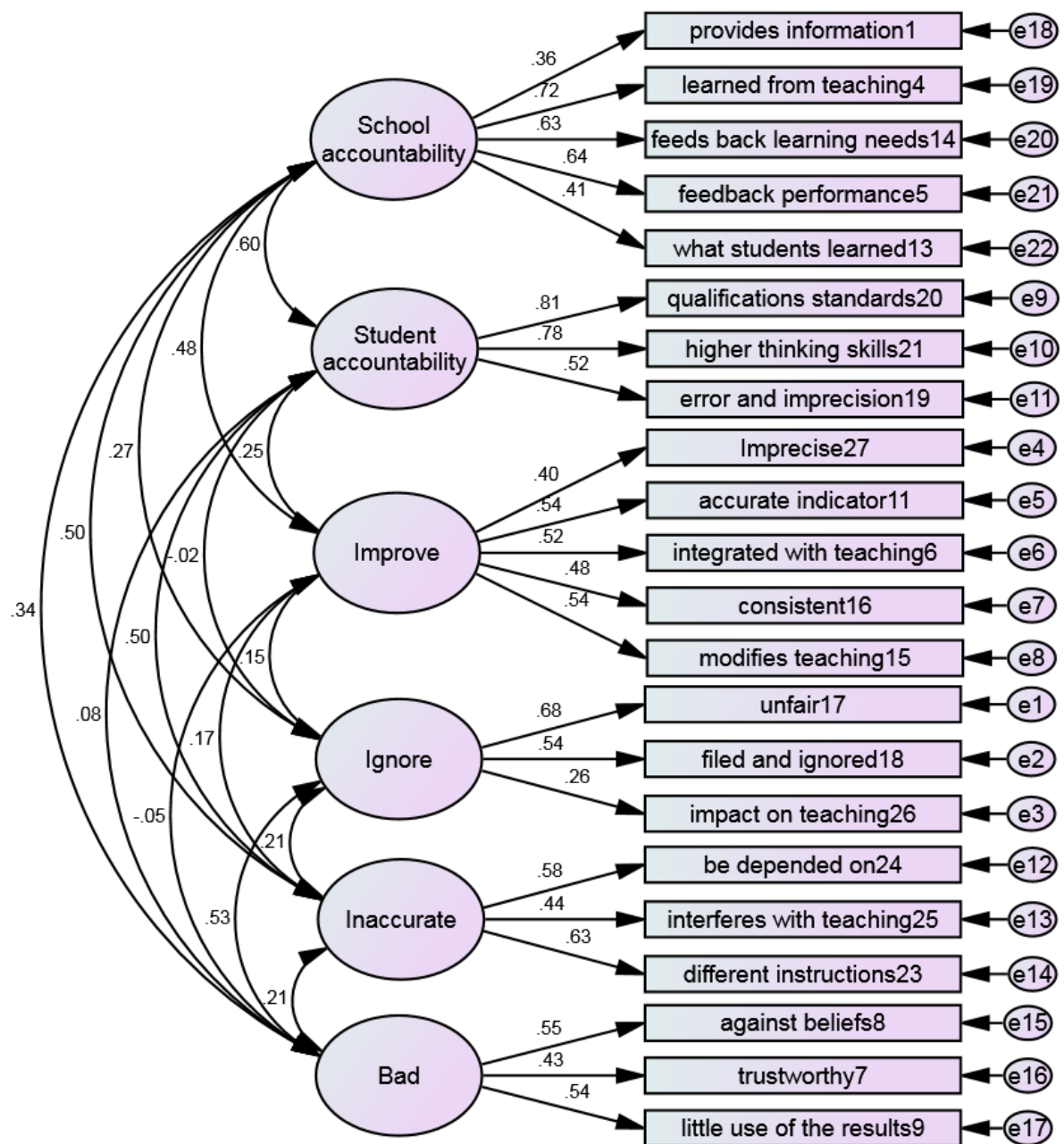
Assessment is an imprecise process.

Appendix B The Tunisian inadmissible model



Fit indices: χ^2 : 1805.788, df : 312, χ^2/df : 5.788, SRMR: .17, RMSEA: .094, CFI: .52, TLI: .46, all loadings were significant at $p = .000$

Appendix C A Tunisian model of conflicting indicators and factors (deselected items, 2,3, 10, 12 and 22)



Fit indices: χ^2 : 525.278, df : 194, χ^2/df : 2.708, SRMR: .10, RMSEA: .056, CFI: .84, TLI: .80, all loadings were significant at $p = .000$

Appendix D Deselected indicators from the data

1. *Assessment is an accurate indicator of a school's quality.*
2. *Assessment establishes what students have learned.*
3. *Assessment is integrated with teaching practice.*
4. *Assessment information modifies ongoing teaching of students.*
5. *Assessment results are trustworthy.*
6. *Assessment results are consistent.*
7. *Assessment results can be depended on.*
8. *Assessment is bad*
9. *Assessment interferes with teaching.*
10. *Assessment results should be treated cautiously given measurement error.*
11. *Teachers should take into account the error and imprecision in all assessment.*
12. *Assessment is an imprecise process.*